

The maximum entropy approach for a posteriori estimation of model and data errors

Svetlana Losa

Alfred Wegener Institute for Polar and Marine Research

Bremerhaven, Germany

Thanks to

Gennady Kivman,
Sergey Danilov,
Frank Janssen,

Jens Schröter,
Vladimir Ryabchenko,
Tijana Janjić

(state x_i and parameter estimation)

Dynamical model

$$L_p(x) = F,$$

L is model operator

uncertainties in initial condition $x(0)$, model parameters p , external forcing F

$$\rho_t^f(x(t), p) = C \rho^f(x(t) | x(0), p) \rho(p) \rho_0(x(0))$$

defined on a $X_t \times P$ space, $x(t) \in X_t, p \in P$

Observational data

$$H(x) = d,$$

H is observational operator

$$\rho(x, p | d) = C \rho(d | x) \rho(x, p)$$

$$\rho(x, p) = C \rho(p) \rho_0(x(0)) \prod_{k=1}^M \rho^f(x(k\delta t) | x((k-1)\delta t), p)$$

We are not confident about the model and data uncertainties.

Do we need the uncertainties quantification?

(general formulation, Kivman et al., 2001)

$$He(\rho) = - \int_X \rho(x|d) \ln \frac{\rho(x|d)}{\mu(x)} \prod dx$$

$\mu(x)$ is the lowest information about x .

The maximum probable x or mean with respect to $\rho(x|d)$ is

$$x_i = M_m x_m + M_d x_d$$

$$M_m = L_* L, \quad M_d = H_* H$$

L_* , H_* reflect our assumptions on the model and data error covariances.

Operators M_m and M_d are nonnegative, self-adjoint and

$$M_m + M_d = I$$

$$(x_i) = \text{Argmin} \left\{ \int_0^T [L(x,t) - F(t)]^2 dt + \beta \sum_{m=1}^M (H(x) - d)^2 \right\}$$

$$He(M) = -\text{trace}(M_d \ln M_d + M_m \ln M_m) = - \sum_{i=1}^N [\lambda_i \ln \lambda_i + (1 - \lambda_i) \ln(1 - \lambda_i)]$$

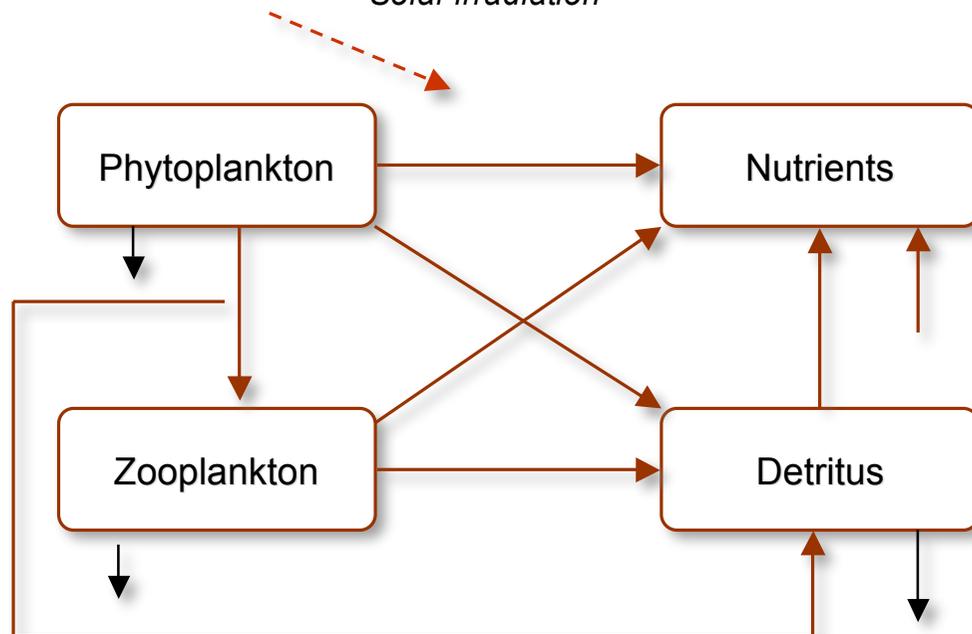
M is an operator-valued measure.

We have to find β which would maximize $He(M)$... or correspondent term in the cost function (Maximum Data Cost, MDC)

Popova's Ecosystem Model (1995)

(generalized inversion)

$$(x, p) = \underset{\text{Solar irradiation}}{\text{Arg min}} \left\{ \int_0^T \left[\frac{dx}{dt} - L_p(x) \right]^2 dt + \beta \sum_{m=1}^M (H(x) - d)^2 \right\}$$



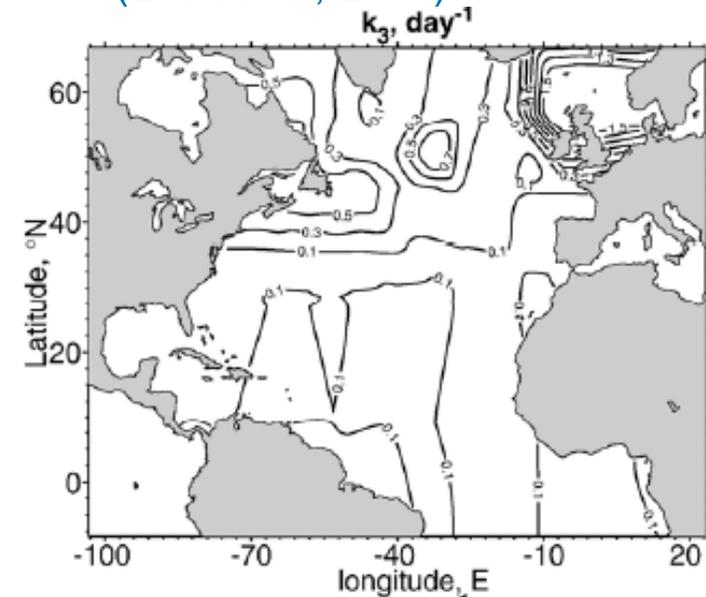
The flow network between 4 biogeochemical {P, Z, N, D} components, x , possesses 19 biological parameters, p .

6 of them have been adjusted for each cell of $5^{\circ} \times 5^{\circ}$ grid covering the North Atlantic

Assimilated data:

Monthly mean satellite CZCS surface chlorophyll averaged over 1979 – 1985.

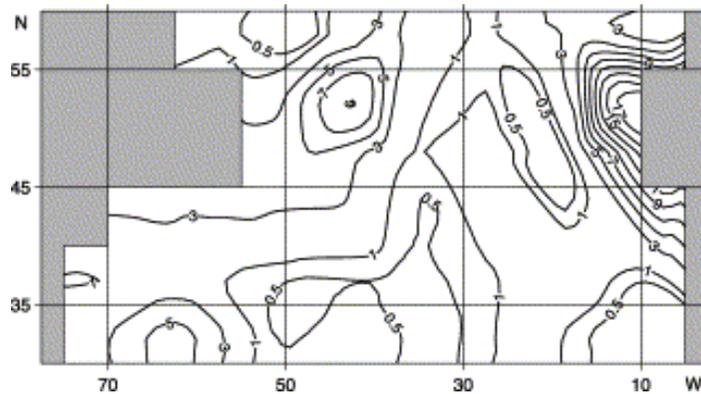
Method : a weak constraint variational technique (Losa et al, 2004)



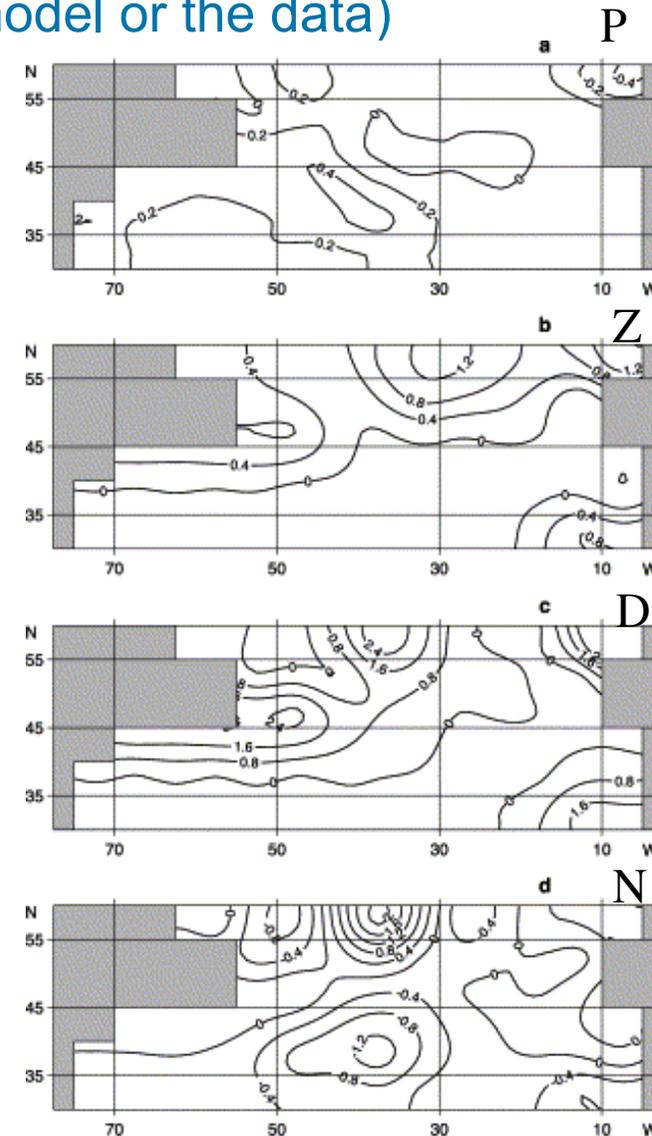
(Which is better: the model or the data)

The ratio of the terms in the cost function

$$r = \frac{\beta \sum_{m=1}^M |H(x) - d|^2}{\int_0^T \left[\frac{dx}{dt} - L_p(x) \right]^2 dt}$$



Annual model equation residuals normalized by the total biological source \Rightarrow



(Secondary Inversion)

$$\frac{d\tilde{x}}{dt} = L_p(x)$$

The first guess

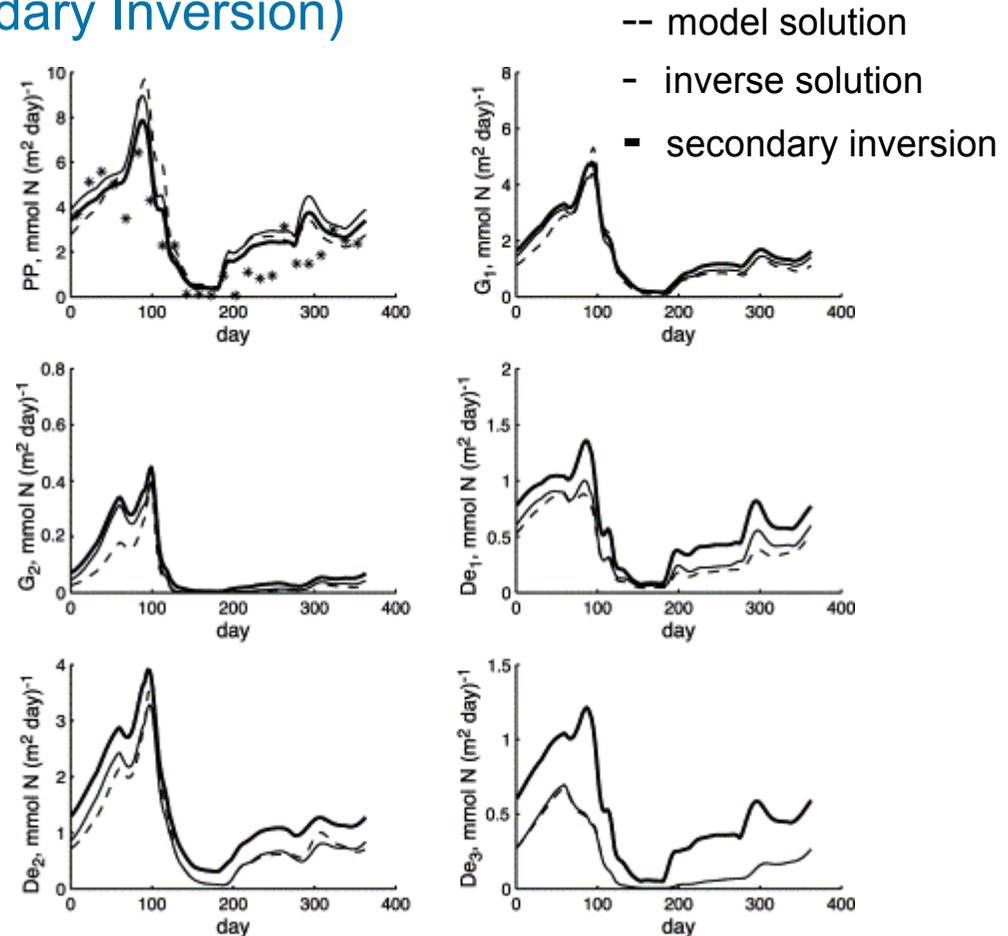
$$P\tilde{P}, \tilde{G}_1, \tilde{G}_2, \tilde{D}e_i, i = 1, 2, 3$$

is calculated given

$$\tilde{x} = (\tilde{P}, \tilde{Z}, \tilde{D}, \tilde{N})$$

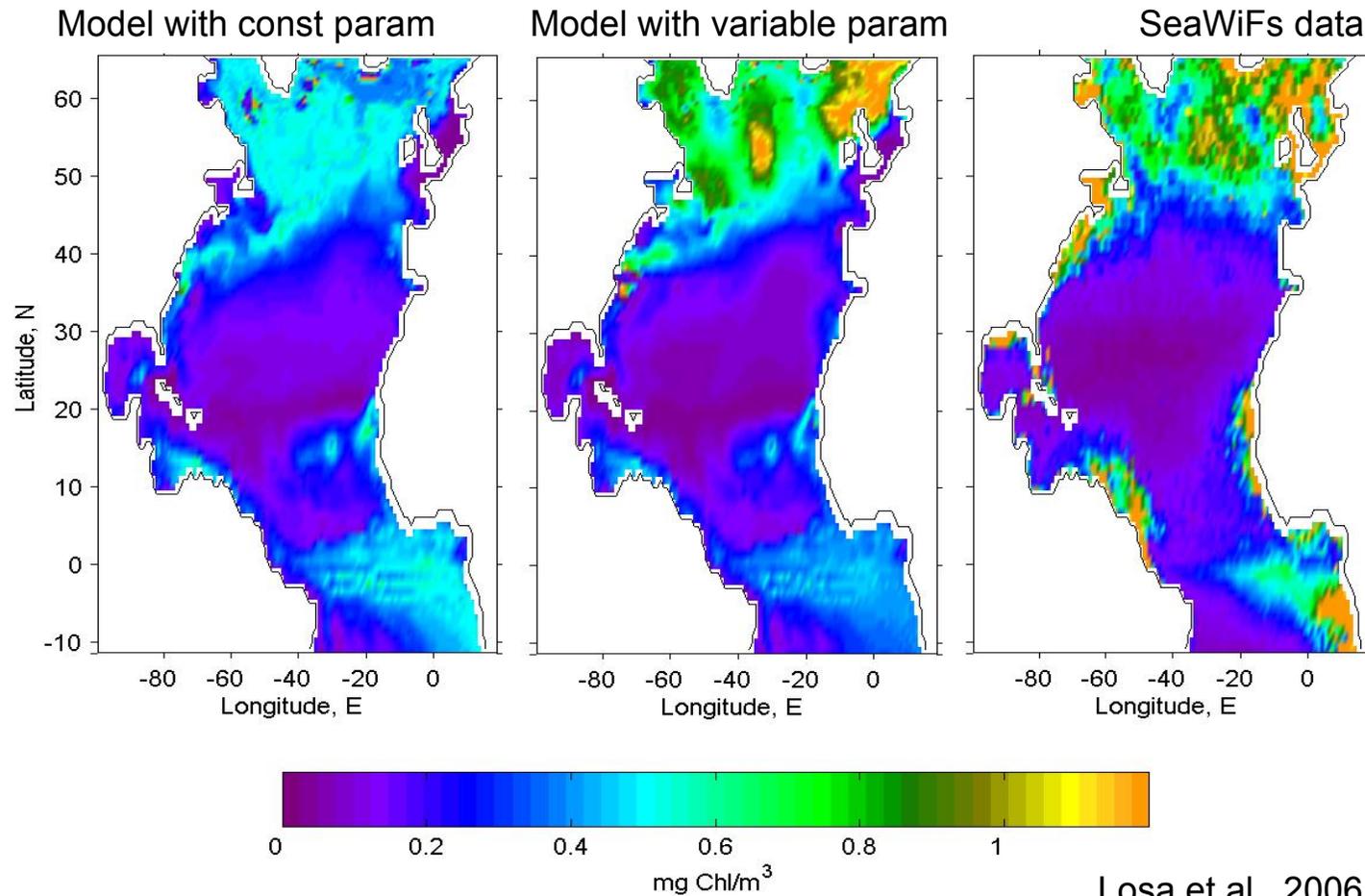
and optimal parameters

$$(PP, G_1, G_2, De_1, De_2, De_3) = \text{Argmin} \left[(PP - P\tilde{P}) + \sum_{i=1}^2 (G_i - \tilde{G}_i)^2 + \sum_{i=1}^3 (De_i - \tilde{D}e_i)^2 \right]$$

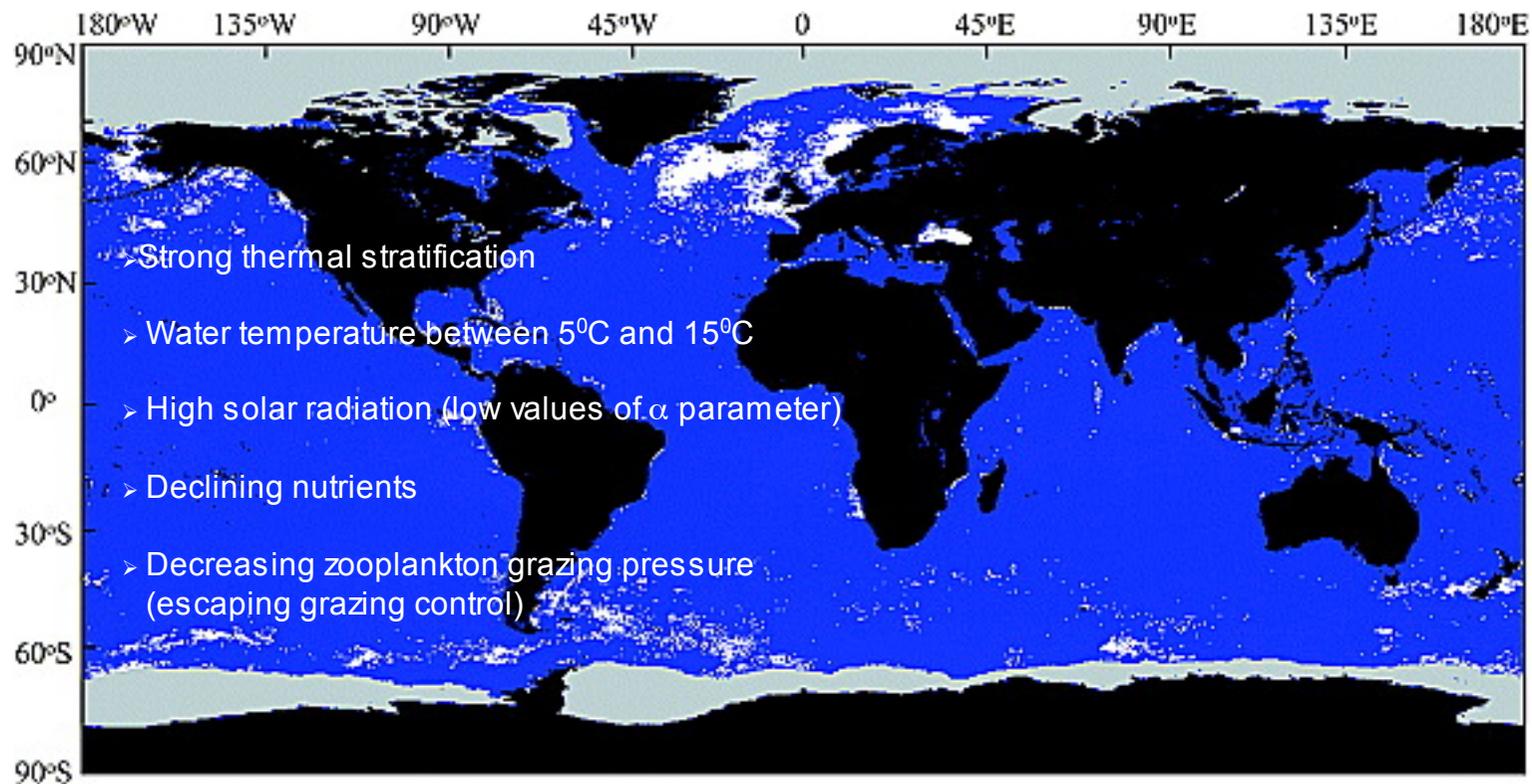


August horizontal distribution of the surface chlorophyll "a" in the North Atlantic

(Popova's NPZD coupled to 3D POP gcm)



Annual composite of classified coccolithophorid blooms in SeaWiFS imagery dating from October 1997 to September 1999 (Iglesias-Rodríguez et al., 2002)



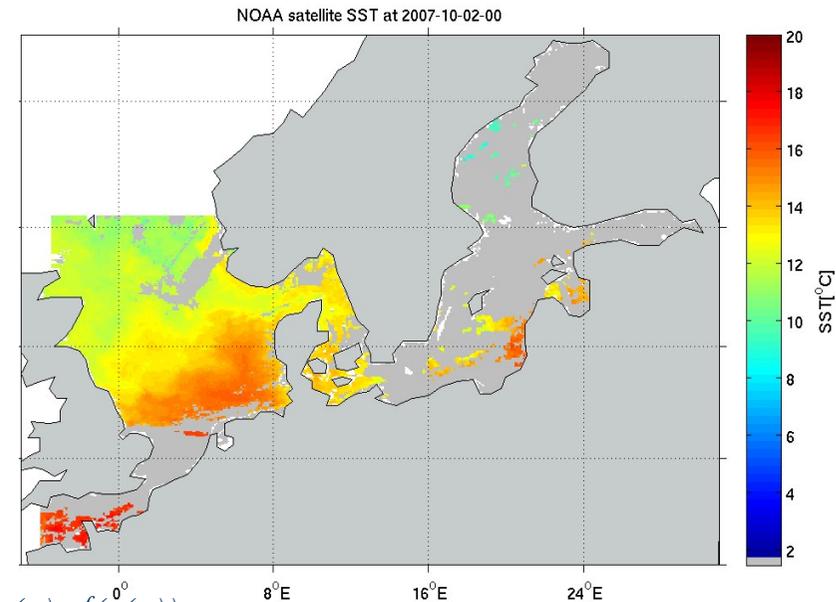
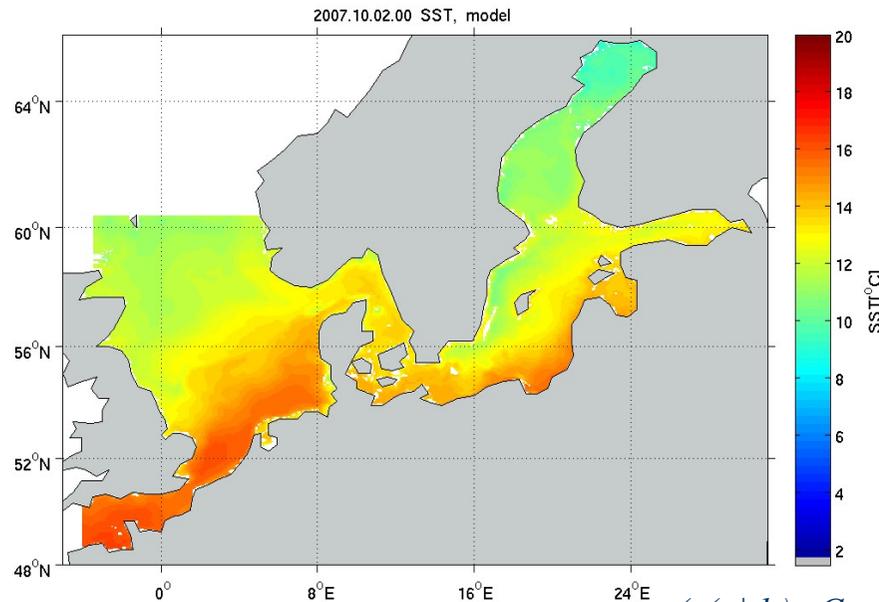
The bloom class is white, the non-coccolithophorid bloom class is blue, the land is black, and ice is gray.

Assimilating NOAA's SST data into an operational circulation model of the North and Baltic Seas

BSHcmod *run at the German Maritime and Hydrographic Agency (BSH)*

NOAA SST

12 hourly-around 00:00 and 12:00,- composites of SST measured by the Advanced Very High Resolution Radiometer (AVHRR) aboard polar orbiting satellites



$$\begin{aligned}\rho_t^a(x(t_1)|d_1) &= C\rho_d(d_1|x(t_1))\rho_t^f(x(t_1)) \\ \rho_t^f(x(t_1)) &= C\rho^f(x(t)|x(0))\rho_0(x(0))\end{aligned}$$

Extraction and combination of the information from two different sources - the model and the data - in order to improve our understanding of both sources and, therefore, of reality itself

(Kalman type filtering)

Ensemble based Singular Evolutive Interpolated Kalman filter (SEIK, Pham, 2001)

$$x(t_n)^a = x(t_n)^{f,m} + K_n (d_n - Hx(t_n)^{f,m})$$

$$K_n = P_n^f H (H P_n^f H^T + R)^{-1}$$

x^f, x^a denote forecast and analysis of state vector (at time t_n at all grid points)

d_n - observations available (at t_n)

P_n^f - forecast error covariance matrix

R - observational error covariance matrix

SEIK Filter is implemented locally (PDAF, Nerger et al., 2006) but with different formulations of data error correlation.

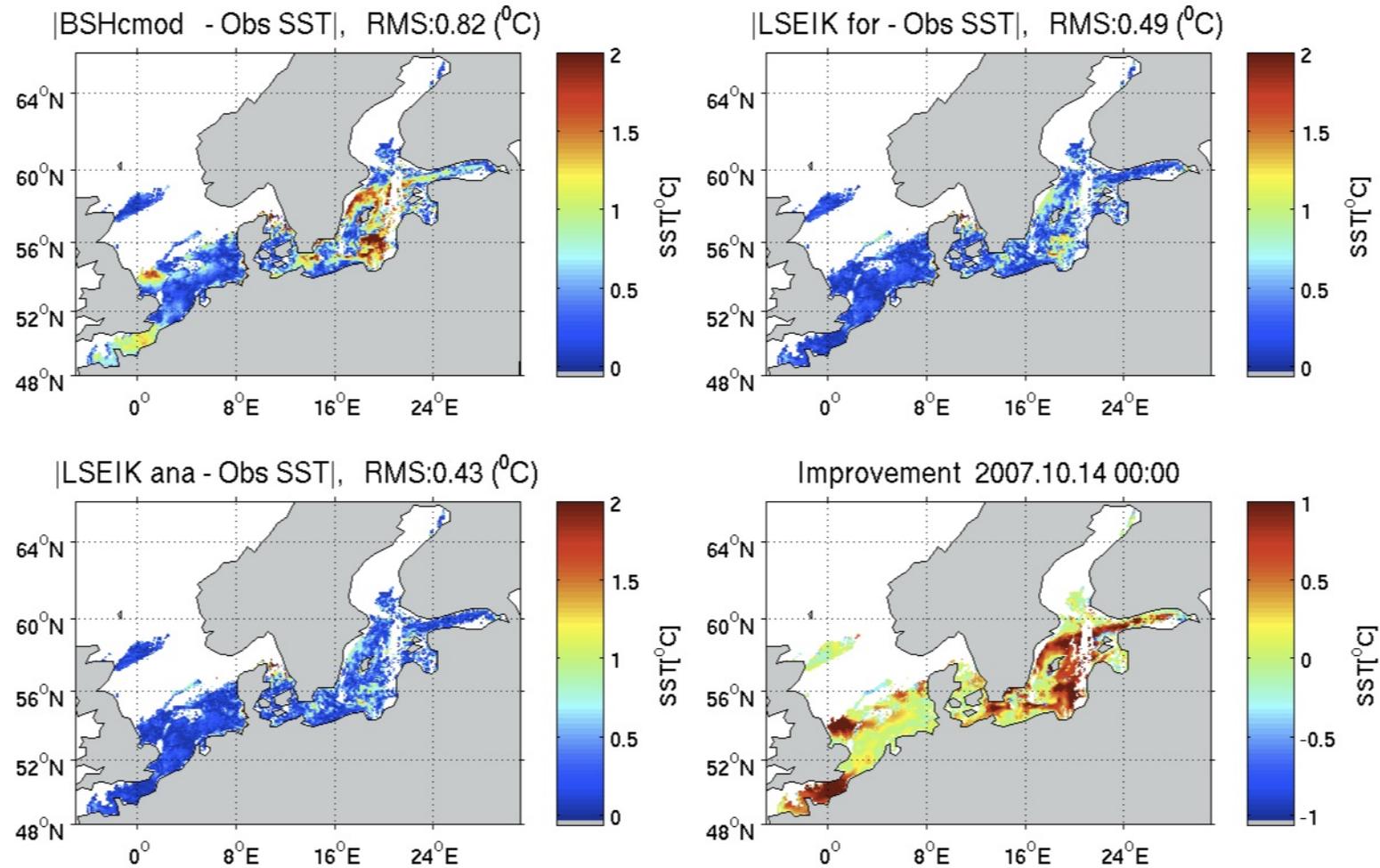
When calculating $He(M)$, the Kalman gain K could be considered

globally over a certain period of time

locally (for validation of localization conditions)

Use SVD decomposition

Improvement of SST analysis and forecast



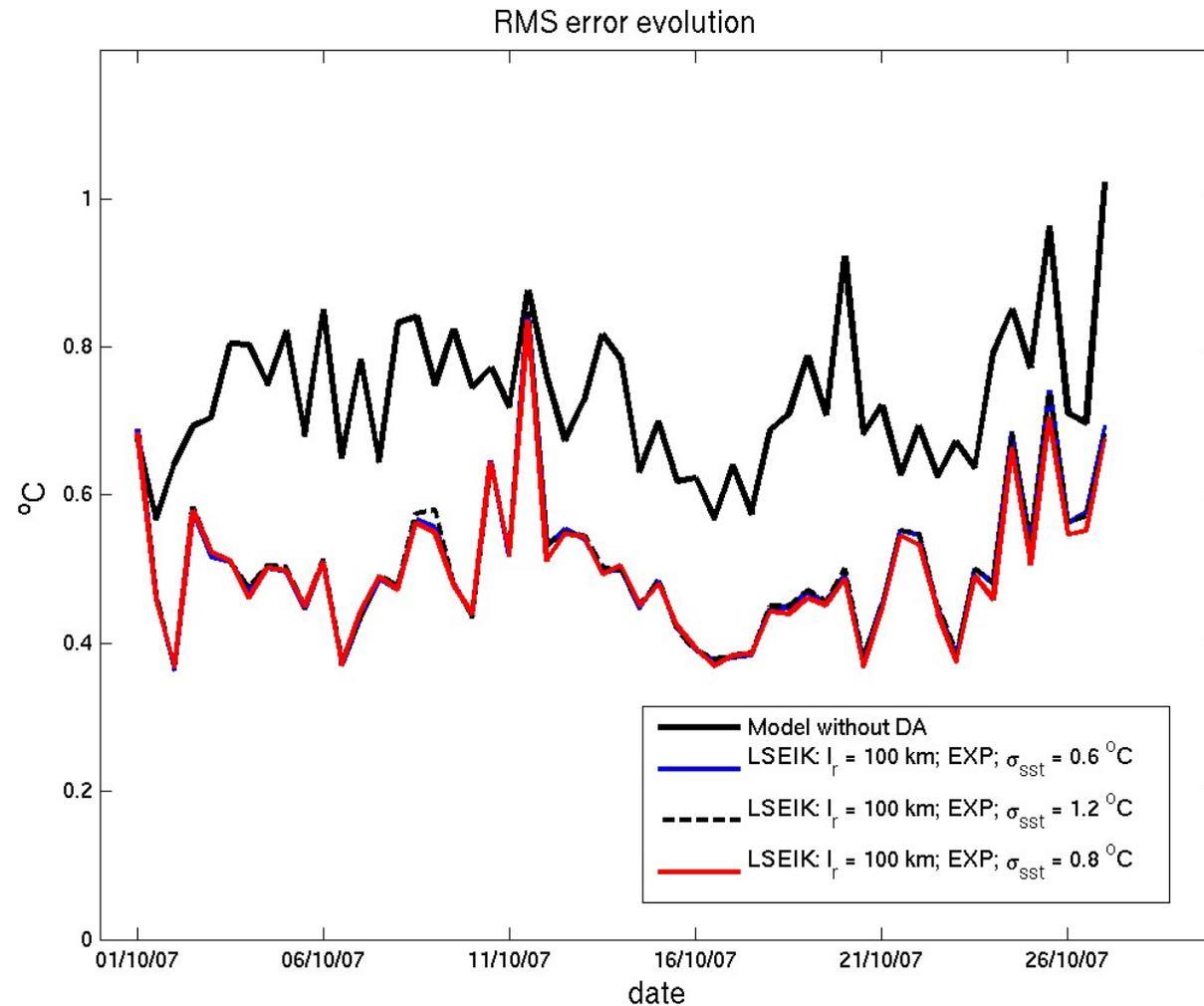
Improvement of SST forecast

Experiment *He(M)*

$\sigma_{sst} = 0.6^\circ\text{C}$ 3.99

$\sigma_{sst} = 0.8^\circ\text{C}$ 4.33

$\sigma_{sst} = 1.2^\circ\text{C}$ 3.90



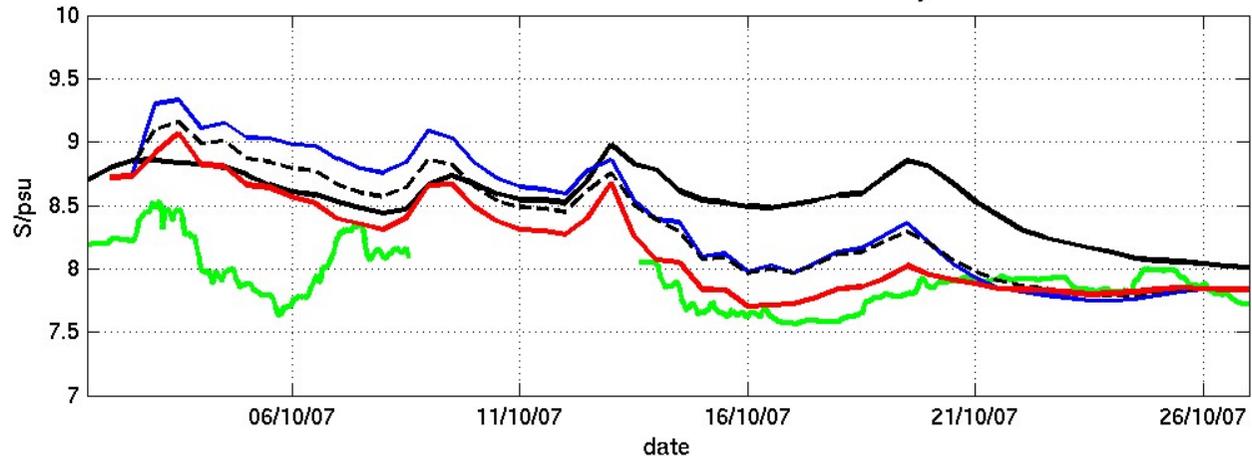
Experiment *He(M)*

$\sigma_{sst} = 0.6^\circ\text{C}$ 3.99

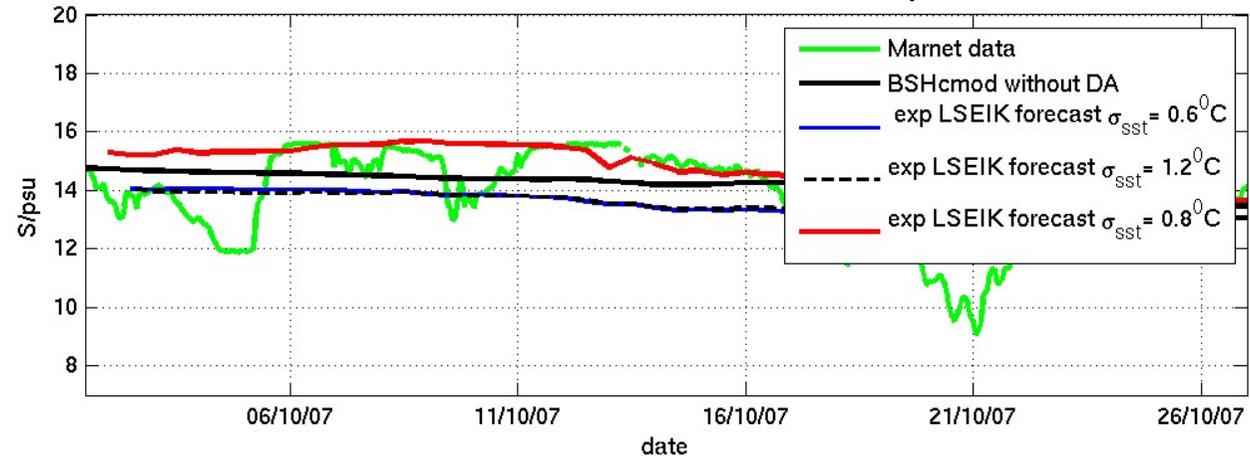
$\sigma_{sst} = 0.8^\circ\text{C}$ 4.33

$\sigma_{sst} = 1.2^\circ\text{C}$ 3.90

Arkona Basin: Surface salinity

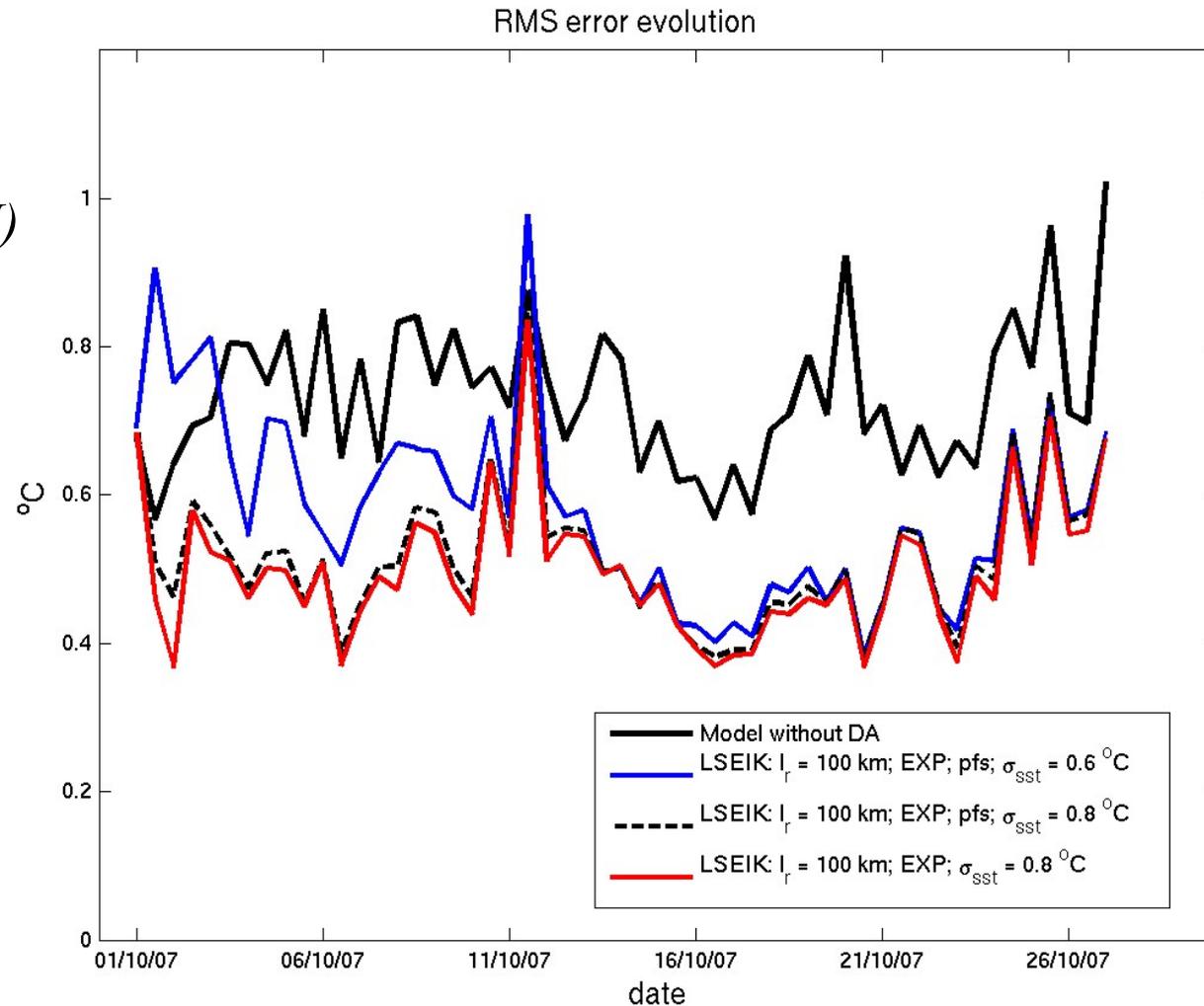


Arkona Basin: Bottom salinity



Sensitivity of the forecast quality

| Experiment | $He(M)$ |
|-----------------------------------|---------|
| $\sigma_{sst} = 0.6^\circ C, Pfs$ | 3.57 |
| $\sigma_{sst} = 0.8^\circ C, Pfs$ | 4.17 |
| $\sigma_{sst} = 0.8^\circ C$ | 4.33 |



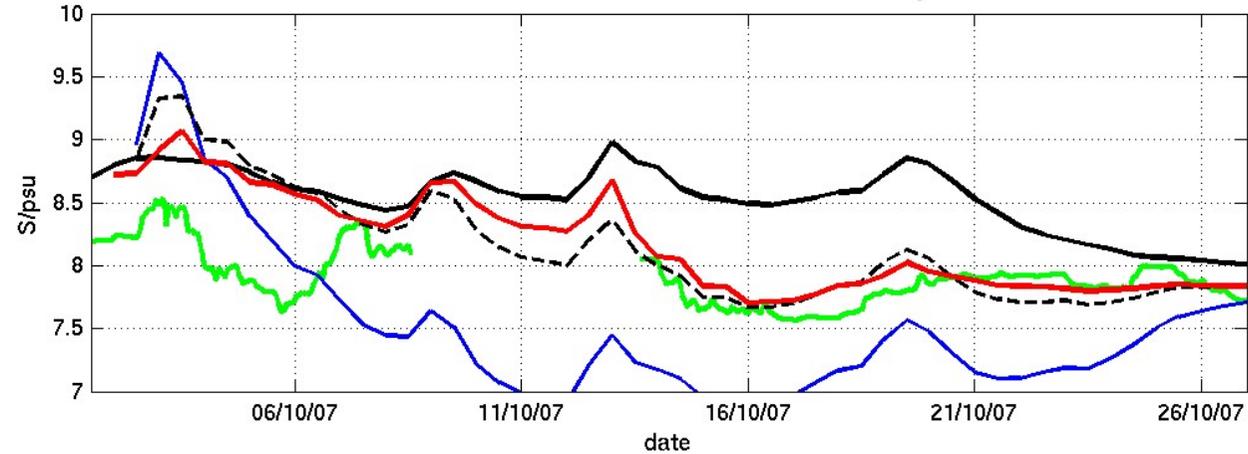
Experiment $He(M)$

$\sigma_{sst} = 0.6^\circ C, Pfs$ 3.57

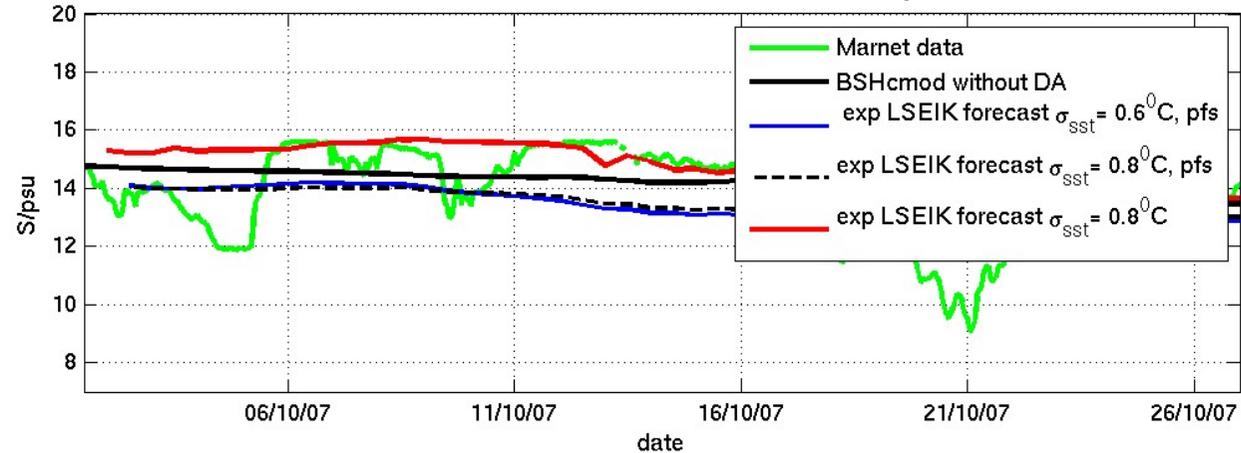
$\sigma_{sst} = 0.8^\circ C, Pfs$ 4.17

$\sigma_{sst} = 0.8^\circ C$ 4.33

Arkona Basin: Surface salinity



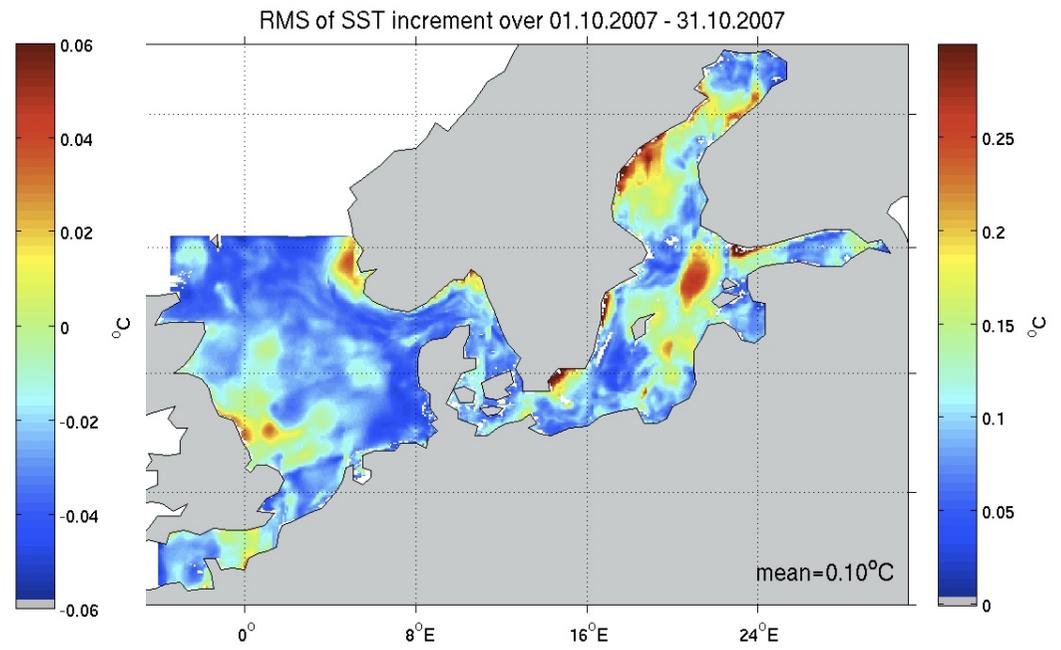
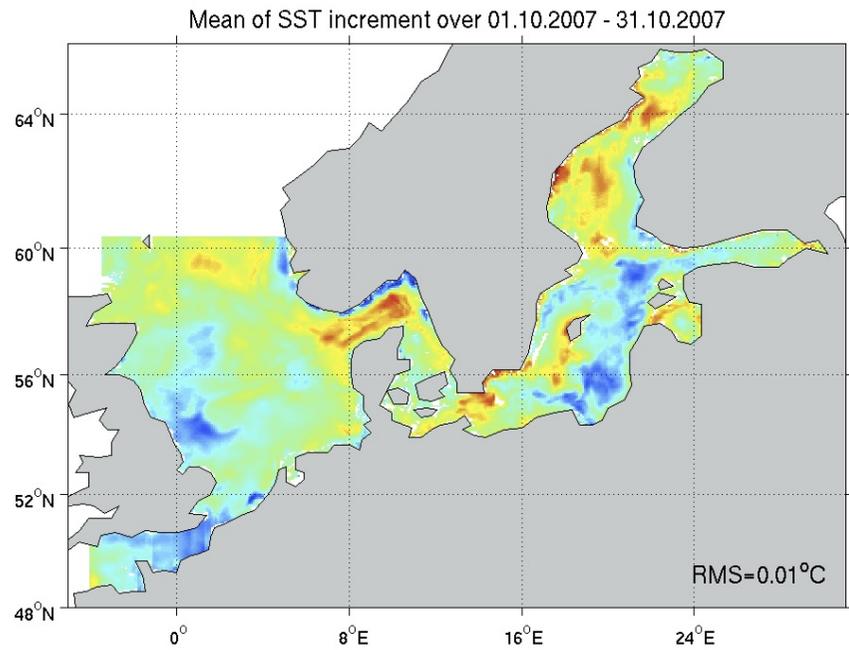
Arkona Basin: Bottom salinity



Deviation from MARNET SST Daten

| Station | RMS (°C) | | | Bias (°C) | | |
|---------|----------|-------|------|-----------|-------|------|
| | Model | LSEIK | NOAA | Model | LSEIK | NOAA |
| Arkona | 0.88 | 0.58 | 0.61 | -0.29 | 0. | 0.04 |
| Darß | 1.27 | 0.81 | 0.69 | -0.55 | -0.17 | 0.01 |
| Kiel | 0.79 | 0.49 | 0.61 | -0.13 | 0.07 | 0.08 |
| Fehm | 0.63 | 0.43 | 0.56 | -0.16 | 0.03 | 0.16 |
| Ems | 0.67 | 0.45 | 0.49 | 0.33 | 0.2 | 0.17 |
| Dbucht | 0.97 | 0.53 | 0.57 | -0.34 | -0.03 | 0.27 |
| nsb | | | 0.73 | | | |

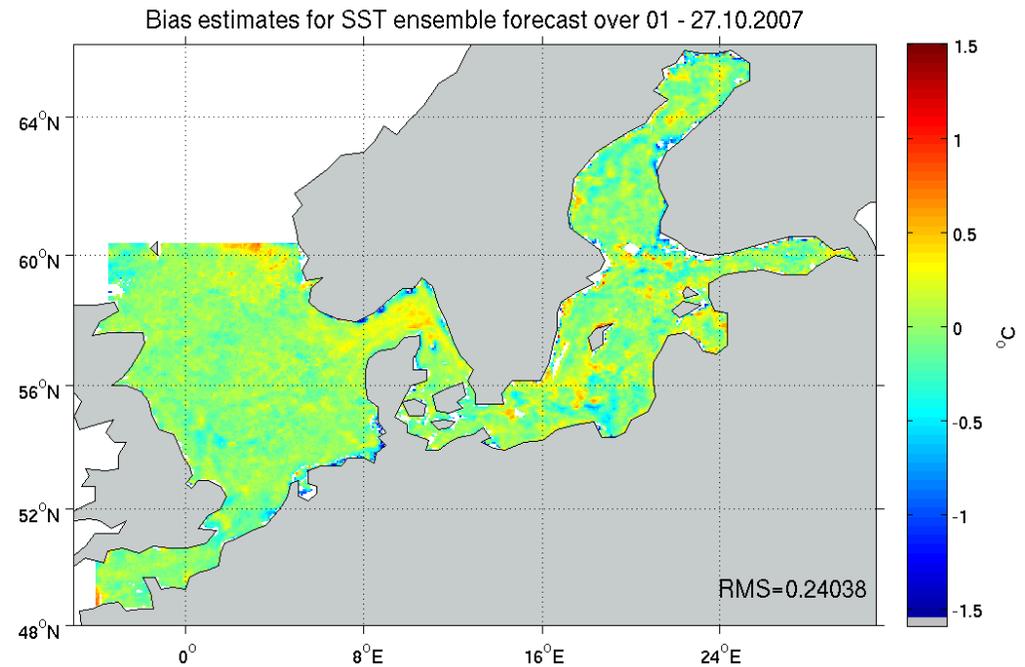
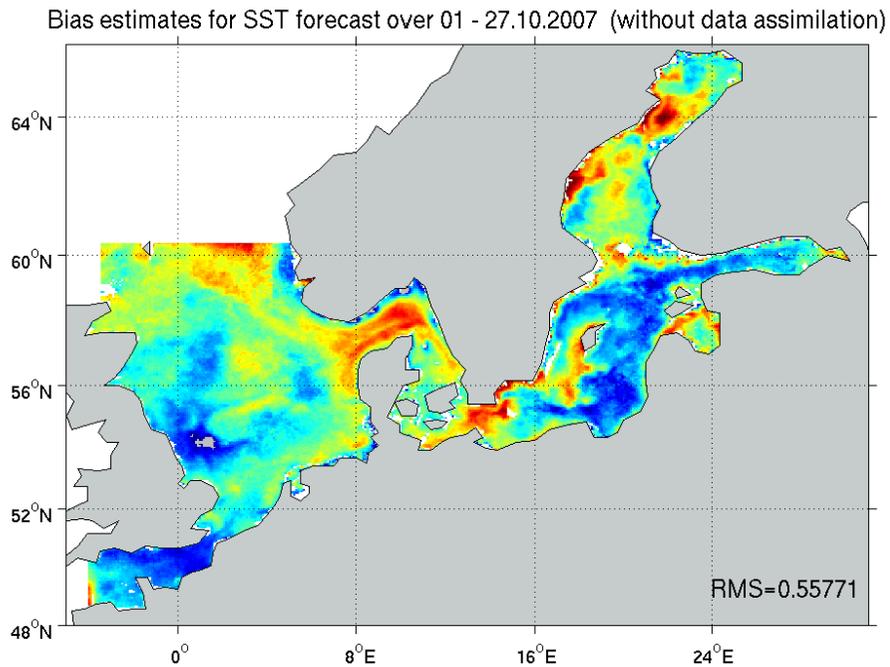
Increment Analysis



Improvement of SST forecast in the North and the Baltic Seas when sequentially assimilating satellite data

Bias without DA

with LSEIK filter



Bias reduction

We have demonstrated two examples of the PME implementation for a posteriori estimating the model and data errors in data assimilation problem.

The chlorophyll satellite data assimilation based on a posteriori choosing of the data weight allowed us to compare the quality of the data and ecosystem model prediction and discern the low quality of the satellite data for high latitudes and for the coastal region of the North Atlantic.

The procedure of the secondary inversion of biogeochemical fluxes makes it possible to restore the mass balance broken while performing the weak constraint parameter estimation and to refine the estimates of the biogeochemical fluxes.

The spatial distribution of the biogeochemical parameters is in a good agreement with independent information about species composition/distribution and their physiology.

Implementation of the PME for assessing prior model and data error statistics in SST data ensemble based assimilation for an operational forecasting model of the North and Baltic Seas revealed the best agreement of the forecast with independent data under the assumptions on initial model and data error statistics, which produced the ME of the posterior distribution.

Investigation of the PME implementation in a local analysis content is of our further interest.

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